NOVEMBER 2004 COURSE 3 SOLUTIONS

Question #1 Key: D

$$E[Z] = \int_{0}^{\infty} b_{t} v^{t}_{t} p_{x} \mu(x+t) dt = \int_{0}^{\infty} e^{0.06t} e^{-0.08t} e^{-0.05t} \frac{1}{20} dt$$

$$= \frac{1}{20} \left(\frac{100}{7} \right) \left[-e^{-0.07t} \right]_{0}^{\infty} = \frac{5}{7}$$

$$E[Z^{2}] = \int_{0}^{\infty} \left(b_{t} v^{t} \right)^{2}_{t} p_{x} \mu(x+t) dt = \int_{0}^{\infty} e^{0.12t} e^{-0.16t} e^{-0.05t} \frac{1}{20} dt = \frac{1}{20} \int_{0}^{\infty} e^{-0.09t} dt$$

$$= \frac{1}{20} \left(\frac{100}{9} \right) \left[e^{-0.09t} \right]_{0}^{\infty} = \frac{5}{9}$$

$$Var[Z] = \frac{5}{9} - \left(\frac{5}{7} \right)^{2} = 0.04535$$

Question #2

Key: C

Let ns = nonsmoker and s = smoker

k =	$q_{x+k}^{(ns)}$	$p_{x+k}^{(ns)}$	$q_{x+k}^{(s)}$	$p_{x+k}^{(s)}$
0	.05	0.95	0.10	0.90
1	.10	0.90	0.20	0.80
2	.15	0.85	0.30	0.70

$$A_{x:\overline{2}|}^{1(ns)} = v \quad q_x^{(ns)} + v^2 \quad p_x^{(ns)} \quad q_{x+1}^{(ns)}$$

$$\frac{1}{1.02} (0.05) \quad \frac{1}{1.02^2} \quad 0.95 \times 0.10 = 0.1403$$

$$A_{x:2}^{1(s)} \qquad v \qquad q_x^{(s)} + v^2 \quad p_x^{(s)} \quad q_{x+1}^{(s)}$$

$$\frac{1}{1.02} \quad (0.10) + \frac{1}{(1.02)^2} \quad 0.90 \times 0.20 = 0.2710$$

$$A_{x:\overline{2}|}^{1}$$
 = weighted average = (0.75)(0.1403) + (0.25)(0.2710)
= 0.1730

Question #3 Key: A

$$\overline{P}(\overline{A}_{x}) = \mu = 0.03$$

$${}^{2}\overline{A}_{x} = 0.20 = \frac{\mu}{2\delta + \mu} = \frac{0.03}{2\delta + 0.03}$$

$$\Rightarrow \delta = 0.06$$

$$\operatorname{Var}({}_{0}L) = \frac{{}^{2}\overline{A}_{x} - (\overline{A}_{x})^{2}}{(\delta \overline{a})^{2}} = \frac{0.20 - (\frac{1}{3})^{2}}{(\frac{0.06}{0.09})^{2}} = 0.20$$
where $A = \frac{\mu}{\mu + \delta} = \frac{0.03}{0.09} = \frac{1}{3}$ $\overline{a} = \frac{1}{\mu + \delta} = \frac{1}{0.09}$

Question #4 Key: B

$$s(60) = \frac{e^{-(0.1)(60)} + e^{-(0.08)(60)}}{2}$$

$$= 0.005354$$

$$s(61) = \frac{e^{-(0.1)(61)} + e^{-(0.08)(61)}}{2}$$

$$= 0.00492$$

$$q_{60} = 1 - \frac{0.00492}{0.005354} = 0.081$$

Key: B

For
$$\Omega$$
, $0.4 = F(\omega) = \left(\frac{\omega}{80}\right)^2$
 $0.6325 = \frac{\omega}{80}$
 $\omega = 50.6$

For
$$T(0)$$
 using De Moivre, $0.7 = F(t) = \frac{t}{\omega} = \frac{t}{50.6}$
 $t = (0.7)(50.6) = 35.42$

Question #6

Key: C

$$E[N] = mq = 1.8 \Rightarrow q = \frac{1.8}{3} = 0.6$$

$$x \qquad f_N(x) \qquad F_N(x)$$

$$0 \qquad 0.064 \qquad 0.064$$

$$1 \qquad 0.288 \qquad 0.352$$

$$2 \qquad 0.432 \qquad 0.784$$

$$3 \qquad 0.216 \qquad 1.000$$

First: 0.432 < 0.7 < 0.784 so N = 2. Use 0.1 and 0.3 for amounts Second: 0.064 < 0.1 < 0.352 so N = 1 Use 0.9 for amount Third: 0.432 < 0.5 < 0.784 so N = 2 Use 0.5 and 0.7 for amounts

Discrete uniform
$$\Rightarrow F_X(x) = 0.2x$$
, $x = 1, 2, 3, 4, 5$
 $0.4 < 0.5 < 0.6 \Rightarrow x_1 = 3$
 $0.6 < 0.7 < 0.8 \Rightarrow x_2 = 4$

Aggregate claims = 3+4=7

Key: C

$$E(X \wedge x) = \frac{\theta}{\alpha - 1} \left[1 - \left(\frac{\theta}{x + \theta} \right)^{\alpha - 1} \right] = \frac{2000x}{x + 2000}$$

$$\begin{array}{c|cc}
x & E(X \wedge x) \\
\hline
\infty & 2000 \\
250 & 222 \\
2250 & 1059 \\
5100 & 1437
\end{array}$$

$$0.75(E(X \land 2250) - E(X \land 250)) + 0.95(E(X) - E(X \land 5100))$$
$$0.75(1059 - 222) + 0.95(2000 - 1437) = 1162.6$$

The 5100 breakpoint was determined by when the insured's share reaches 3600:

$$3600 = 250 + 0.25(2250 - 250) + (5100 - 2250)$$

Question #8

Key: D

Since each time the probability of a heavy scientist is just half the probability of a success, the distribution is binomial with $q = 0.6 \times 0.5 = 0.3$ and m = 8.

$$f(2) = (8 \times 7/2) \times (0.3^2) \times (0.7^6) = 0.30$$

Key: A

$$\mu_{xy}(t) = \mu_x(t) + \mu_y(t) = 0.08 + 0.04 = 0.12$$

$$\overline{A}_x = \mu_x(t) / (\mu_x(t) + \delta) = 0.5714$$

$$\overline{A}_y = \mu_y(t) / (\mu_y(t) + \delta) = 0.4$$

$$\overline{A}_{xy} = \mu_{xy}(t) / (\mu_{xy}(t) + \delta) = 0.6667$$

$$\overline{a}_{xy} = 1 / (\mu_{xy}(t) + \delta) = 5.556$$

$$\overline{A}_{xy} = \overline{A}_x + \overline{A}_y - \overline{A}_{xy} = 0.5714 + 0.4 - 0.6667 = 0.3047$$
Premium = 0.304762/5.556 = 0.0549

Question #10 Key: B

$$P_{40} = A_{40} / \ddot{a}_{40} = 0.16132 / 14.8166 = 0.0108878$$

$$P_{42} = A_{42} / \ddot{a}_{42} = 0.17636 / 14.5510 = 0.0121201$$

$$a_{45} = \ddot{a}_{45} - 1 = 13.1121$$

$$E \left[{}_{3}L \middle| K (42) \ge 3 \right] = 1000 A_{45} - 1000 P_{40} - 1000 P_{42} a_{45}$$

$$= 201.20 - 10.89 - (12.12)(13.1121)$$

$$= 31.39$$

Many similar formulas would work equally well. One possibility would be $1000_3V_{42} + (1000P_{42} - 1000P_{40})$, because prospectively after duration 3, this differs from the normal benefit reserve in that in the next year you collect $1000P_{40}$ instead of $1000P_{42}$.

Key: E

$$\dot{e}_{x} = \frac{\omega - x}{2}$$

$$k|q_{x} = \frac{1}{\omega - x}$$

$$A_{x} = \sum_{k=b}^{\omega - x-1} v^{k+1} k|q_{x} = \frac{1}{\omega - x} \sum_{k=b}^{\omega - x-1} v^{k+1}$$

$$A_{x} = \frac{a_{\omega - x}|}{\omega - x}$$

$$\ddot{a}_{x} = \frac{1 - A_{x}}{d}$$

$$\stackrel{\circ}{e}_{50} = 25 \Longrightarrow \omega = 100$$
 for typical annuitants

$$\stackrel{\circ}{e}_y = 15 \Rightarrow y = \text{Assumed age} = 70$$

$$A_{70} = \frac{a_{\overline{30}|}}{30} = 0.45883$$

$$\ddot{a}_{70} = 9.5607$$

$$500000 = b \ \ddot{a}_{20} \Rightarrow b = 52,297$$

Question #12 Key: D

$$p_x^{(\tau)} = p_x^{(1)} p_x^{(2)} = 0.8(0.7) = 0.56$$

$$q_x^{(1)} = \left[\frac{\ln \left(p_x^{(1)} \right)}{\ln \left(p_x^{(\tau)} \right)} \right] q_x^{(\tau)} \text{ since UDD in double decrement table}$$

$$= \left[\frac{\ln (0.8)}{\ln (0.56)} \right] 0.44$$

$$= 0.1693$$

$$0.3 q_{x+0.1}^{(1)} = \frac{0.3 q_x^{(1)}}{1 - 0.1 q_x^{(\tau)}} = 0.053$$

To elaborate on the last step:

$$q_{x+0.1}^{(1)} = \frac{\left(\text{Number dying from cause} \atop 1 \text{ between } x + 0.1 \text{ and } x + 0.4\right)}{\text{Number alive at } x + 0.1}$$

Since UDD in double decrement,

$$= \frac{l_x^{(\tau)}(0.3)q_x^{(1)}}{l_x^{(\tau)}(1-0.1q_x^{(\tau)})}$$

Key: B

non absorbing matrix $T = \begin{pmatrix} 0.7 & 0.1 \\ 0.3 & 0.6 \end{pmatrix}$, the submatrix excluding "Terminated", which is an absorbing state.

$$I - T = \begin{pmatrix} 0.3 & -0.1 \\ -0.3 & 0.4 \end{pmatrix}$$
$$(I - T)^{-1} = \begin{pmatrix} \frac{0.4}{0.09} & \frac{0.1}{0.09} \\ \frac{0.3}{0.09} & \frac{0.3}{0.09} \end{pmatrix} = \begin{pmatrix} 4.4\overline{4} & 1.1\overline{1} \\ 3.3\overline{3} & 3.3\overline{3} \end{pmatrix}$$

Future costs for a healthy = $4.4\overline{4} \times 500 + 1.1\overline{1} \times 3000$ = 5555

Question #14

Key: D

$$T = \begin{pmatrix} 0.7 & 0.1 & 0.2 \\ 0.3 & 0.6 & 0.1 \\ 0 & 0 & 1 \end{pmatrix} \qquad T^2 = \begin{pmatrix} 0.52 & 0.13 & 0.35 \\ 0.39 & 0.39 & 0.22 \\ 0 & 0 & 1 \end{pmatrix}$$

Actuarial present value (A.P.V.) prem = 800(1 + (0.7 + 0.1) + (0.52 + 0.13)) = 1,960A.P.V. claim = 500(1 + 0.7 + 0.52) + 3000(0 + 0.1 + 0.13) = 1800Difference = 160

Key: A

Let N_1 , N_2 denote the random variable for # of claims for Type I and II in 2 years

 X_1 , X_2 denote the claim amount for Type I and II

 S_1 = total claim amount for type I in 2 years

 S_2 = total claim amount for Type II at time in 2 years

 $S = S_1 + S_2 = \text{total claim amount in 2 years}$

$$\{S_1\} \rightarrow \text{compound poisson } \lambda_1 = 2 \times 6 = 12 \quad X_1 \sim U(0,1)$$

$$\{S_2\} \rightarrow \text{compound poisson } \lambda_2 = 2 \times 2 = 4 \quad X_2 \sim U(0, 5)$$

$$E(N_1) = Var(N_1) = 2 \times 6 = 12$$

$$E(S_1) = E(N_1)E(X_1) = (12)(0.5) = 6$$

$$Var(S_1) = E(N_1)Var(X_1) + Var(N_1)(E(X_1))^2$$
$$= (12)\frac{(1-0)}{12} + (12)(0.5)^2$$

$$E(N_2) = Var(N_2) = 2 \times 2 = 4$$

With formulas corresponding to those for S_1 ,

$$E(S_2) = 4 \times \frac{5}{2} = 10$$

$$Var(S_2) = 4 \times \frac{(5-0)^2}{12} + 4\left(\frac{5}{2}\right)^2 = 33.\overline{3}$$

$$E(S) = E(S_1) + E(S_2) = 6 + 10 = 16$$

Since S_1 and S_2 are independent,

$$Var(S) = Var(S_1) + Var(S_2) = 4 + 33.\overline{3} = 37.\overline{3}$$

$$\Pr(S > 18) = \Pr\left(\frac{S - 16}{\sqrt{39.3}} > \frac{2}{\sqrt{37.3}} = 0.327\right)$$

Using normal approximation

$$Pr(S > 18) = 1 - \Phi(0.327)$$

$$=0.37$$

Key: D

Since the rate of depletion is constant there are only 2 ways the reservoir can be empty sometime within the next 10 days.

Way #1:

There is no rainfall within the next 5 days

Way #2

There is one rainfall in the next 5 days

And it is a normal rainfall

And there are no further rainfalls for the next five days

Prob (Way #1) = Prob(0 in 5 days) =
$$\exp(-0.2*5) = 0.3679$$

Prob (Way #2) = Prob(1 in 5 days) × 0.8 × Prob(0 in 5 days)
= $5*0.2 \exp(-0.2*5)*0.8 * \exp(-0.2*5)$
= $1 \exp(-1) * 0.8 * \exp(-1) = 0.1083$

Hence Prob empty at some time = 0.3679 + 0.1083 = 0.476

Question #17

Key: C

Let *X* be the loss random variable,

So (X-5) is the claim random variable.

$$E(X) = \frac{10}{2.5 - 1} = 6.\overline{6}$$

$$E(X \land 5) = \left(\frac{10}{2.5 - 1}\right) \left[1 - \left(\frac{10}{5 + 10}\right)^{2.5 - 1}\right]$$

$$= 3.038$$

$$E(X-5)_{+} = E(X) - E(X \land 5)$$
$$= 6.\overline{6} - 3.038$$
$$= 3.629$$

Expected aggregate claims =
$$E(N)E(X-5)_{+}$$

= $(5)(3.629)$
= 18.15

Key: B

A Pareto
$$(\alpha = 2, \theta = 5)$$
 distribution with 20% inflation becomes Pareto with $\alpha = 2, \ \theta = 5 \times 1.2 = 6$
In 2004, $E(X) = \frac{6}{2-1} = 6$
 $E(X \wedge 10) = \frac{6}{2-1} \left(1 - \left(\frac{6}{10+6}\right)^{2-1}\right) = 3.75$
 $E(X-10)_{+} = E(X) - E(X \wedge 10)$
 $= 6 - 3.75 = 2.25$
LER = $1 - \frac{E(X-10)_{+}}{E(X)} = 1 - \frac{2.25}{6} = 0.625$

Question #19

Key: A

Let
$$X =$$
 annual claims
 $E(X) = (0.75)(3) + (0.15)(5) + (0.1)(7)$
 $= 3.7$
 $\pi =$ Premium $= (3.7)(\frac{4}{3}) = 4.93$
Change during year $= 4.93 - 3 = +1.93$ with $p = 0.75$
 $= 4.93 - 5 = -0.07$ with $p = 0.15$
 $= 4.93 - 7 = -2.07$ with $p = 0.10$

Since we start year 1 with surplus of 3, at end of year 1 we have 4.93, 2.93, or 0.93 (with associated probabilities 0.75, 0.15, 0.10).

We cannot drop more then 2.07 in year 2, so ruin occurs only if we are at 0.93 after 1 and have a drop of 2.07.

$$Prob = (0.1)(0.1) = 0.01$$

Key: E

$$0.96 = e^{-(\mu_1 + \lambda)}$$

$$\mu_1 + \lambda = -\ln(0.96) = 0.04082$$

$$\mu_1 = 0.04082 - \lambda = 0.04082 - 0.01 = 0.03082$$

Similarly

$$\begin{split} \mu_2 &= -\ln(0.97) - \lambda = 0.03046 - 0.01 = 0.02046 \\ \mu_{xy} &= \mu_1 + \mu_2 + \lambda = 0.03082 + 0.02046 + 0.01 = 0.06128 \\ {}_5p_{xy} &= e^{-(5)(0.06128)} = e^{-0.3064} = 0.736 \end{split}$$

Key: C

$$A_{60} = 0.36913$$
 $d = 0.05660$
 $^{2}A_{60} = 0.17741$
and $\sqrt{^{2}A_{60} - A_{60}^{2}} = 0.202862$

Expected Loss on one policy is
$$E[L(\pi)] = (100,000 + \frac{\pi}{d})A_{60} - \frac{\pi}{d}$$

Variance on one policy is
$$\operatorname{Var}\left[L(\pi)\right] = \left(100,000 + \frac{\pi}{d}\right)^2 \left({}^2A_{60} - A_{60}^2\right)$$

On the 10000 lives,

$$E[S] = 10,000E[L(\pi)]$$
 and $Var[S] = 10,000 Var[L(\pi)]$

The π is such that $0 - E[S] / \sqrt{Var[S]} = 2.326$ since $\Phi(2.326) = 0.99$

$$\frac{10,000\left(\frac{\pi}{d} - \left(100,000 + \frac{\pi}{d}\right)A_{60}\right)}{100\left(100,000 + \frac{\pi}{d}\right)\sqrt{{}^{2}A_{60} - A_{60}^{2}}} = 2.326$$

$$\frac{100\left(\frac{\pi}{d} - \left(100,000 + \frac{\pi}{d}\right)\right)(0.36913)}{\left(100,000 + \frac{\pi}{d}\right)(0.202862)} = 2.326$$

$$\frac{0.63087\frac{\pi}{d} - 36913}{100,000 + \frac{\pi}{d}} = 0.004719$$

$$0.63087 \frac{\pi}{d} - 36913 = 471.9 = 0.004719 \frac{\pi}{d}$$

$$\frac{\pi}{d} = \frac{36913 + 471.9}{0.63087 - 0.004719}$$
$$= 59706$$

$$\pi = 59706 \times d = 3379$$

Question #22 Key: C

$$_{1}V = (_{0}V + \pi)(1+i) - (1000 + _{1}V - _{1}V) \times q_{75}$$

= 1.05 π - 1000 q_{75}

Similarly,

$${}_{2}V = \binom{1}{1}V + \pi) \times 1.05 - 1000q_{76}$$

$${}_{3}V = \binom{1}{2}V + \pi) \times 1.05 - 1000q_{77}$$

$$1000 = {}_{3}V = \left(1.05^{3}\pi + 1.05^{2} \cdot \pi + 1.05\pi\right) - 1000 \times q_{75} \times 1.05^{2} - 1000 \times 1.05 \times q_{76} - 1000 \times q_{77} \quad *$$

$$\pi = \frac{1000 + 1000\left(1.05^{2}q_{75} + 1.05q_{76} + q_{77}\right)}{\left(1.05\right)^{3} + \left(1.05\right)^{2} + 1.05}$$

$$= \frac{1000^{x}\left(1 + 1.05^{2} \times 0.05169 + 1.05 \times 0.05647 + 0.06168\right)}{3.310125}$$

$$= \frac{1000 \times 1.17796}{3.310125} = 355.87$$

* This equation is algebraic manipulation of the three equations in three unknowns $({}_1V, {}_2V, \pi)$. One method – usually effective in problems where benefit = stated amount plus reserve, is to multiply the ${}_1V$ equation by 1.05^2 , the ${}_2V$ equation by 1.05, and add those two to the ${}_3V$ equation: in the result, you can cancel out the ${}_1V$, and ${}_2V$ terms. Or you can substitute the ${}_1V$ equation into the ${}_2V$ equation, giving ${}_2V$ in terms of π , and then substitute that into the ${}_3V$ equation.

Question #23 Key: D

Actuarial present value (APV) of future benefits =
$$= (0.005 \times 2000 + 0.04 \times 1000) / 1.06 + (1 - 0.005 - 0.04) (0.008 \times 2000 + 0.06 \times 1000) / 1.06^{2}$$

$$= 47.17 + 64.60$$

$$= 111.77$$
APV of future premiums =
$$[1 + (1 - 0.005 - 0.04) / 1.06] 50$$

$$= (1.9009)(50)$$

= 95.05
$$E \begin{bmatrix} _1L | K(55) \ge 1 \end{bmatrix}$$
 = 111.77 - 95.05 = 16.72

Question #24 Key: D

Alternate approach: if losses are Pareto with $\theta = 50$ and $\alpha = 3$, then claim payments per payment with an ordinary deductible of 20 are Pareto with $\theta = 50 + 20$ and $\alpha = 3$.

Thus
$$E(T(20)) = \frac{50 + 20}{3 - 1} = 35$$

This alternate approach was shown here for educational reasons: to reinforce the idea that many life contingent models and non-life models can have similar structure. We doubt many candidates would take that approach, especially since it involves specific properties of the Pareto distribution.

Key: E

 $Q \ge P$ since in Q you only test at intervals; surplus below 0 might recover before the next test. In P, ruin occurs if you are ever below 0.

 $R \ge P$ since you are less likely to have surplus below 0 in the first N years (finite horizon) than forever.

Add the inequalities

$$Q + R \ge 2P$$

Also (why other choices are wrong)

 $S \ge Q$ by reasoning comparable to $R \ge P$. Same testing frequency in S and Q, but Q tests forever.

 $S \ge R$ by reasoning comparable to $Q \ge P$. Same horizon in S and R; R tests more frequently.

 $S \ge P$ P tests more frequently, and tests forever.

Question #26

Key: A

This is a nonhomogeneous Poisson process with intensity function

$$\lambda$$
 (t) = 3+3t, $0 \le t \le 2$, where t is time after noon

Average
$$\lambda = \frac{\int_{1}^{2} \lambda(t)dt}{1} = \int_{1}^{2} (3+3t)dt$$
$$= \left[3t + \frac{3t^{2}}{2}\right]_{1}^{2}$$
$$= 7.5$$

$$f(2) = \frac{e^{-7.5} 7.5^2}{2!} = 0.0156$$

Key: E

X(t) - Y(t) is Brownian motion with initial value -2 and $\sigma^2 = 0.5 + 1 = 1.5$

By formula 10.6, the probability that $X(t) - Y(t) \ge 0$ at some time between 0 and 5 is

$$= 2 \times \operatorname{Prob}\left[X(5) - Y(5) \ge 0\right]$$
$$= 2 \times \left[1 - \Phi\left(\frac{2}{\sqrt{5(1.5)}}\right)\right] = 0.4652$$

The 2 in the numerator of $\frac{2}{\sqrt{5(1.5)}}$ comes from X(0)-Y(0)=-2; the process needs to move 2 to reach $X(t)-Y(t) \ge 0$.

Key: B

$${}_{2|}q_{\overline{80:84}} = {}_{2|}q_{80} + {}_{2|}q_{84} - {}_{2|}q_{80:84}$$
$$= 0.5 \times 0.4 \times (1 - 0.6) + 0.2 \times 0.15 \times (1 - 0.1)$$
$$= 0.10136$$

Using new p_{82} value of 0.3

$$0.5 \times 0.4 \times (1 - 0.3) + 0.2 \times 0.15 \times (1 - 0.1)$$

= 0.16118

Change =
$$0.16118 - 0.10136 = 0.06$$

Alternatively,

$${}_{2}p_{80} = 0.5 \times 0.4 = 0.20$$

$${}_{3}p_{80} = {}_{2}p_{80} \times 0.6 = 0.12$$

$${}_{2}p_{84} = 0.20 \times 0.15 = 0.03$$

$${}_{3}p_{84} = {}_{2}p_{84} \times 0.10 = 0.003$$

$${}_{2}p_{\overline{80:84}} = {}_{2}p_{80} + {}_{2}p_{84} - {}_{2}p_{80} \cdot {}_{2}p_{84} \text{ since independent}$$

$$= 0.20 + 0.03 - (0.20)(0.03) = 0.224$$

$${}_{3}p_{\overline{80:84}} = {}_{3}p_{80} + {}_{3}p_{84} - {}_{3}p_{80} \cdot {}_{3}p_{84}$$

$$= 0.12 + 0.003 - (0.12)(0.003) = 0.12264$$

$${}_{2}|q_{\overline{80:84}} = {}_{2}p_{\overline{80:84}} - {}_{3}p_{\overline{80:84}}$$

$$= 0.224 - 0.12264 = 0.10136$$

Revised

$${}_{3}p_{80} = 0.20 \times 0.30 = 0.06$$

$${}_{3}p_{\overline{80:84}} = 0.06 + 0.003 - (0.06)(0.003)$$

$$= 0.06282$$

$${}_{2|}q_{\overline{80:84}} = 0.224 - 0.06282 = 0.16118$$
change = 0.16118 - 0.10136 = 0.06

Question #29 Key: B

$$e_x = p_x + p_x e_{x+1} \Rightarrow p_x = \frac{e_x}{1 + e_{x+1}} = \frac{8.83}{9.29} = 0.95048$$

$$\ddot{a}_x = 1 + v p_x + v_2^2 p_x + \dots$$

$$\ddot{a}_{\overline{x.2}} = 1 + v + v_2^2 p_x + \dots$$

$$\ddot{a}_{x.\overline{2}} - \ddot{a}_x = v q_x = 5.6459 - 5.60 = 0.0459$$

$$v (1 - 0.95048) = 0.0459$$

$$v = 0.9269$$

$$i = \frac{1}{v} - 1 = 0.0789$$

Key: A

Let π be the benefit premium

Let $_{\iota}V$ denote the benefit reserve a the end of year k.

For any
$$n$$
, $\binom{n}{v} + \pi (1+i) = (q_{25+n} \times_{n+1} V + p_{25+n} \times_{n+1} V)$
= $\binom{n}{v} = \binom{n}{v} + \pi (1+i)$

Thus
$$_{1}V = (_{0}V + \pi)(1+i)$$

$${}_{2}V = ({}_{1}V + \pi)(1+i) = (\pi(1+i) + \pi)(1+i) = \pi \ddot{s}_{\overline{2}|}$$

$${}_{3}V = ({}_{2}V + \pi)(1+i) = (\pi \ddot{s}_{\overline{2}|} + \pi)(1+i) = \pi \ddot{s}_{\overline{3}|}$$

By induction (proof omitted)

$$_{n}V=\pi\ddot{s}_{\overline{n}}$$

For n = 35, ${}_{n}V = \ddot{a}_{60}$ (actuarial present value of future benefits; there are no future premiums)

$$\ddot{a}_{60} = \pi \ \ddot{s}_{\overline{35}}$$

$$\pi = \frac{\ddot{a}_{60}}{\ddot{s}_{\overline{35}|}}$$

For
$$n = 20$$
, ${}_{20}V = \pi \ddot{s}_{\overline{20}|}$
$$= \left(\frac{\ddot{a}_{60}}{\ddot{s}_{2\overline{5}|}}\right) \ddot{s}_{\overline{20}|}$$

Alternatively, as above

$$\binom{n}{V} + \pi (1+i) = \binom{n}{n+1}V$$

Write those equations, for n = 0 to n = 34

$$0:(_{0}V+\pi)(1+i)=_{1}V$$

$$1:({}_{1}V+\pi)(1+i)={}_{2}V$$

$$2: \left({}_{2}V + \pi\right)\left(1+i\right) = {}_{3}V$$

$$34:(_{34}V+\pi)(1+i)=_{35}V$$

Multiply equation k by $(1+i)^{34-k}$ and sum the results:

$$(_{0}V + \pi)(1+i)^{35} + (_{1}V + \pi)(1+i)^{34} + (_{2}V + \pi)(1+i)^{33} + \dots + (_{34}V + \pi)(1+i) =$$

$${_{1}V(1+i)^{34}} + {_{2}V(1+i)^{33}} + {_{3}V(1+i)^{32}} + \dots + {_{34}V(1+i)} + {_{35}V}$$

For
$$k = 1, 2, \dots, 34$$
, the $_kV (1+i)^{35-k}$ terms in both sides cancel, leaving $_0V (1+i)^{35} + \pi \Big[(1+i)^{35} + (1+i)^{34} + \dots + (1+i) \Big] =_{35}V$
Since $_0V = 0$
 $\pi \ddot{s}_{\overline{35}} =_{35}V$
 $= \ddot{a}_{60}$

(see above for remainder of solution)

This technique, for situations where the death benefit is a specified amount (here, 0) plus the benefit reserve is discussed in section 8.3 of Bowers. This specific problem is Example 8.3.1.

Key: B

$$\mu_{\overline{xy}}(t) = \frac{{}_{t}q_{y} {}_{t}p_{x}\mu(x+t) + {}_{t}q_{x} {}_{t}p_{y}\mu(y+t)}{{}_{t}q_{x} \times {}_{t}p_{y} + {}_{t}p_{x} \times {}_{t}q_{y} + {}_{t}p_{x} \times {}_{t}p_{y}}$$

For
$$(x) = (y) = (50)$$

$$\mu_{\overline{50:50}}(10.5) = \frac{\binom{10.5}{q_{50}}\binom{1}{10}p_{50}q_{60} \cdot 2}{\binom{10.5}{q_{50}}\binom{1}{10.5}p_{50} \cdot 2 + \binom{10.5}{10.5}p_{50}^{2}} = \frac{(0.09152)(0.91478)(0.01376)(2)}{(0.09152)(0.90848)(2) + (0.90848)^{2}} = 0.0023$$

where

$$_{10.5}p_{50} = \frac{\frac{1/2(l_{60} + l_{61})}{l_{50}}}{l_{50}} = \frac{\frac{1/2(8,188,074 + 8,075,403)}{8,950,901} = 0.90848$$

$$q_{50} = 1 - q_{50} = 0.09152$$

$$_{10} p_{50} = \frac{8,188,074}{8,950,901} = 0.91478$$

$$_{10.5} p_{50} \mu (50+10.5) = (_{10} p_{50}) q_{60}$$
 since UDD

Alternatively,
$$_{(10+t)} p_{50} = _{10} p_{50} t p_{60}$$

$$_{(10+t)}p_{50:50} = (_{10}p_{50})^2 (_t p_{60})^2$$

$$(10+t) p_{\overline{50:50}} = 2 {}_{10} p_{50} {}_{t} p_{60} - ({}_{10} p_{50})^{2} ({}_{t} p_{60})^{2}$$

$$= 2 {}_{10} p_{50} (1 - tq_{60}) - ({}_{10} p_{50})^{2} (1 - tq_{60})^{2} \text{ since UDD}$$

Derivative =
$$-2_{10} p_{50} q_{60} + 2(_{10} p_{50})^2 (1 - tq_{60}) q_{60}$$

Derivative at 10 + t = 10.5 is

$$-2(0.91478)(0.01376) + (0.91478)^{2}(1 - (0.5)(0.01376))(0.01376) = -0.0023$$

$$p_{10.5} p_{\overline{50:50}} = 2_{10.5} p_{50} - (1_{10.5} p_{50})^{2}$$
$$= 2(0.90848) - (0.90848)^{2}$$
$$= 0.99162$$

$$\mu$$
 (for any sort of lifetime) = $\frac{-\frac{dp}{dt}}{p} = \frac{-(-0.0023)}{0.99162} = 0.0023$

Key: E

$$E(W) = \frac{1}{4} \int_{0}^{4} \sum_{i=0}^{\infty} 2^{i} \Pr(N = i | \lambda) d\lambda \qquad \left[\frac{1}{4} \text{ is the density of } \lambda \text{ on } [0, 4]. \right]$$

$$= \frac{1}{4} \int_{0}^{4} P(2 | \lambda) d\lambda \qquad \text{[see note]}$$

$$= \frac{1}{4} \int_{0}^{4} e^{\lambda(2-1)} d\lambda \qquad \text{[using formula from tables for the pgf of the Poisson]}$$

$$= \frac{1}{4} e^{\lambda} \Big|_{0}^{4} = \frac{1}{4} (e^{4} - 1)$$

$$= 13.4$$

Note: the probability generating function (pgf) is $P(Z) = \sum_{k=0}^{\infty} p_k Z^k$ so the integrand is P(2), or in this case $P(2|\lambda)$ since λ is not known.

Alternatively,

$$E(W) = \frac{1}{4} \int_{0}^{4} \sum_{i=0}^{\infty} 2^{i} \Pr(N = i | \lambda) d\lambda$$

$$= \frac{1}{4} \int_{0}^{4} \sum_{i=0}^{\infty} \frac{2^{i} e^{-\lambda} \lambda^{i}}{i!} d\lambda$$

$$= \frac{1}{4} \int_{0}^{4} \sum_{i=0}^{\infty} \frac{e^{-\lambda} (2\lambda)^{i}}{i!} d\lambda$$
We know
$$\sum_{i=0}^{\infty} \frac{e^{-2\lambda} (2\lambda)^{i}}{i!} = 1 \text{ since } \frac{e^{-2\lambda} (2\lambda)^{i}}{i!} \text{ is } f(i) \text{ for a Poisson with mean } Z\lambda$$

so
$$\sum_{i=0}^{\infty} \frac{e^{-\lambda} (2\lambda)^i}{i!} = \frac{e^{-\lambda}}{e^{-2\lambda}} = e^{\lambda}$$

Thus
$$E(W) = \frac{1}{4} \int_0^4 e^{\lambda} d\lambda$$

$$= \frac{1}{4}e^{\lambda} \Big|_{0}^{4} = \frac{1}{4}(e^{4} - 1)$$

$$= 13.4$$

Question #33 Key: A or E

$$E(S) = \lambda E[X] = 2/3(1/4 + 2/4 + 3/2) = 2/3 \times 9/4 = 3/2$$

 $Var(S) = \lambda E[X^2] = 2/3(1/4 + 4/4 + 9/2) = 23/6$

So cumulative premium to time 2 is $2(3/2+1.8\sqrt{23/6})=10$, where the expression in parentheses is the annual premium

Times between claims are determined by $-1/\lambda \log u$ and are 0.43, 0.77, 1.37, 2.41 So 2 claims before time 2 (second claim is at 1.20; third is at 2.57)

Sizes are 2, 3, 1, 3, where only the first two matter.

So gain to the insurer is 10-(2+3) = 5

Note: since the problem did not specify that we wanted the gain or loss from the insurer's viewpoint, we gave credit to answer A; a loss of 5 from the insured's viewpoint.

Key: C

To get number of claims, set up cdf for Poisson:

X	f(x)	F(x)
0	0.135	0.135
1	0.271	0.406
2	0.271	0.677
3	0.180	0.857

0.80 simulates 3 claims.

$$F(x) = 1 - (500/(x+500))^2 = u$$
, so $x = (1-u)^{-1/2} 500 - 500$

0.6 simulates 290.57

0.25 simulates 77.35

0.7 simulates 412.87

So total losses equals 780.79

Insurer pays (0.80)(750) + (780.79 - 750) = 631

Question #35 Key: D

$$\mu_{x}^{(\tau)}(t) = \mu_{x}^{(1)}(t) + \mu_{x}^{(2)}(t) = 0.01 + 2.29 = 2.30$$

$$P = P \int_{0}^{2} v_{t}^{t} p_{x}^{(\tau)} \mu_{x}^{(2)}(t) dt + 50,000 \int_{0}^{2} v_{t}^{t} p_{x}^{(\tau)} \mu_{x}^{(1)}(t) dt + 50,000 \int_{2}^{\infty} v_{t}^{t} p_{x}^{(\tau)} \mu_{x}^{(\tau)}(t) dt$$

$$P = P \int_{0}^{2} e^{-0.1t} e^{-2.3t} \times 2.29 dt + 50,000 \int_{0}^{2} e^{-0.1t} e^{-2.3t} \times 0.01 dt + 50,000 \int_{2}^{\infty} e^{-0.1t} e^{-2.3t} \times 2.3 dt$$

$$P \left[1 - 2.29 \times \frac{1 - e^{-2(2.4)}}{2.4} \right] = 50000 \left[0.01 \times \frac{1 - e^{-2(2.4)}}{2.4} + 2.3 \times \frac{e^{-2(2.4)}}{2.4} \right]$$

$$P = 11,194$$

Question #36 Key: D

$$\mu^{(accid)} = 0.001$$

$$\mu^{(total)} = 0.01$$

$$\mu^{(other)} = 0.01 - 0.001 = 0.009$$

Actuarial present value =
$$\int_0^\infty 500,000 e^{-0.05t} e^{-0.01t} (0.009) dt$$

 $+10 \int_0^\infty 50,000 e^{0.04t} e^{-0.05t} e^{-0.01t} (0.001) dt$
 $= 500,000 \left[\frac{0.009}{0.06} + \frac{0.001}{0.02} \right] = 100,000$

Key: B

Variance =
$$v^{30}_{15}p_{x\,15}q_x$$
 Expected value = $v^{15}_{15}p_x$
 $v^{30}_{15}p_{x\,15}q_x = 0.065$ $v^{15}_{15}p_x$
 $v^{15}_{15}q_x = 0.065 \Rightarrow_{15}q_x = 0.3157$

Since μ is constant

$$q_x = \left(1 - (p_x)^{15}\right)$$
$$(p_x)^{15} = 0.6843$$
$$p_x = 0.975$$
$$q_x = 0.025$$

Question #38

Key: E

$$(1) \quad {}_{11}V^A = \left({}_{10}V^A + 0\right) \frac{(1+i)}{p_{x+10}} - \frac{q_{x+10}}{p_{x+10}} \times 1000$$

(2)
$$_{11}V^B = (_{10}V^B + \pi^B)\frac{(1+i)}{p_{x+10}} - \frac{q_{x+10}}{p_{x+10}} \times 1000$$

$$(1)-(2) _{11}V^{A} - _{11}V^{B} = \left(_{10}V^{A} - _{10}V^{B} - \pi^{B}\right)\frac{(1+i)}{p_{x+10}}$$
$$= (101.35 - 8.36)\frac{(1.06)}{1 - 0.004}$$
$$= 98.97$$

Key: A

Actuarial present value Benefits =
$$\frac{(0.8)(0.1)(10,000)}{1.06^2} + \frac{(0.8)(0.9)(0.097)(9,000)}{1.06^3}$$
$$= 1,239.75$$

$$1,239.75 = P\left(1 + \frac{(0.8)}{1.06} + \frac{(0.8)(0.9)}{1.06^2}\right)$$
$$= P(2.3955)$$
$$P = 517.53 \implies 518$$

Question #40

Key: C

Event
$$x = 0$$
 Prob Present Value 15
 $x = 1$ $(0.95)(0.10) = 0.095$ $15 + 20/1.06 = 33.87$
 $x \ge 2$ $(0.95)(0.90) = 0.855$ $15 + 20/1.06 + 25/1.06^2 = 56.12$

$$E[X] = (0.05)(15) + (0.095)(33.87) + (0.855)(56.12) = 51.95$$

$$E[X^{2}] = (0.05)(15)^{2} + (0.095)(33.87)^{2} + (0.855)(56.12)^{2} = 2813.01$$

$$Var[X] = E(X^{2}) - E(X)^{2} = 2813.01 - (51.95)^{2} = 114.2$$